# Exact BER Analysis for Alamouti's Code on Arbitrary Fading Channels with Imperfect Channel Estimation 

Lennert Jacobs and Marc Moeneclaey<br>Ghent University, TELIN Department, DIGCOM Group<br>Sint-Pietersnieuwstraat 41, B-9000 Gent, Belgium<br>\{Lennert.Jacobs,Marc.Moeneclaey\} @ telin.ugent.be


#### Abstract

In this contribution, we investigate the effect of imperfect channel estimation on the bit error rate (BER) performance of multiple-input multiple-output (MIMO) systems employing Alamouti's code. The propagation channels from both transmit antennas to each of the $N_{\mathrm{r}}$ receive antennas are assumed to be affected by (possibly correlated) flat block fading with an arbitrary fading distribution. The transmitted symbols belong to a pulse amplitude modulation (PAM) or quadrature amplitude modulation (QAM) constellation. The mismatched maximumlikelihood (ML) receiver makes use of a least-squares or linear minimum mean-square error (MMSE) channel estimate, obtained from known pilot symbols sent among the data.

The resulting average BER for QAM transmission can easily be written as an expectation over $8 N_{\mathrm{r}}$ real-valued random variables, but the computing time needed for its numerical evaluation increases dramatically with $N_{\mathrm{r}}$. We point out that the BER can be expressed in terms of the distribution of the Frobenius norm of the channel matrix, rather than the joint distribution of all channel coefficients. This allows to reduce the BER expression for any number of receive antennas to an expectation over only 6 random variables, which can easily be computed numerically. Moreover, we show that for real-valued constellations and/or realvalued channels, the BER expression reduces to an expectation over less than 6 variables. For BER levels of practical interest, the numerical evaluation of the BER is much less time-consuming than a straightforward computer simulation. The presented BER expression is useful not only when the fading distribution is given in closed form, but also when only experimental data (e.g. a histogram) on the fading are available.


## I. Introduction

The performance of wireless communication systems is strongly affected by small-scale fading. To combat this detrimental phenomenon, several diversity techniques have been proposed to provide the receiver with independent replicas of the signal. Spatial diversity is achieved by using multiple antennas at the transmitter and/or receiver side. These socalled multiple-input multiple-output (MIMO) communication systems can achieve a maximum diversity order of $N_{\mathrm{t}} N_{\mathrm{r}}$ (with $N_{\mathrm{t}}$ and $N_{\mathrm{r}}$ denoting the number of transmit and receive antennas, respectively), provided that proper space-time coding is used. In 1998, Alamouti invented a simple coding scheme for data transmission using two transmit antennas [1], with the remarkable benefit that the maximum-likelihood (ML) decoding algorithm reduces to symbol-by-symbol detection, based only on linear processing at the receiver. Tarokh et al. generalized Alamouti's scheme to an arbitrary number of
transmit antennas by introducing the concept of orthogonal space-time block coding [17], [18]. As the orthogonal spacetime block codes (OSTBCs) achieve full spatial diversity, and require only linear processing at the receiver, these codes are a very attractive transmit diversity technique. Under the assumption of perfect channel knowledge (PCK), the bit error rate (BER) performance of OSTBCs has been studied extensively in e.g. [4], [9], [10] and [11].
In practical wireless applications, however, the receiver has to estimate the channel response, which inevitably results in a performance penalty. Most often, the investigation of the resulting performance has been carried out under the assumption of Rayleigh fading: an analytical expression for the BER of OSTBCs in case of minimum mean-square error (MMSE) channel estimation was derived in [8]; in [15], analytical BER expressions as well as the tight Chernoff bound were given for orthogonal space-time block coded systems employing $M$-ary phase-shift keying ( $M$-PSK) modulation; high-SNR expressions for the pairwise error probability (PEP) were derived in [2] for quite general STBCs with coherent and noncoherent receivers, using an eigenvalue approach; in [5], an exact closed-form expression for the PEP of both orthogonal and non-orthogonal space-time codes in the case of least-squares channel estimation was obtained by means of characteristic functions. However, from the PEP one can compute only an upper bound on the BER, which in a fading environment does not converge to the true BER at high SNR. In [6], expressions for the exact decoding error probability (DEP) were presented for the case of square OSTBCs on arbitrary fading channels, but the analysis is restricted to PSK constellations.
In this contribution, we provide an exact BER analysis for Alamouti's code with pulse amplitude modulation (PAM) or quadrature amplitude modulation (QAM) constellations on arbitrary flat-fading channels with imperfect channel estimation (ICE). In section II we describe the observation model which includes the $2 N_{\mathrm{r}}$ arbitrary fading channels. Section III presents a linear pilot-based channel estimation method (with the wellknown least-squares estimation and linear MMSE estimation as special cases), and derives the statistical properties of the channel estimate. The mismatched ML receiver is briefly outlined in section IV. The main part of our contribution is in section V , where the exact BER expression is reduced
to an expectation over (at most) 6 variables, that allows numerical evaluation with a computing time that is independent of the number ( $N_{r}$ ) of receive antennas. In section VI the results obtained from the numerical evaluation of our BER expression are confirmed by computer simulations, assuming i.i.d. Nakagami- $m$ fading channels. Finally, conclusions are drawn in section VII. The main conclusion is that the presented BER expression can be used when the fading distribution is available in closed form or as an experimentally obtained histogram, and allows for a faster evaluation than by means of straightforward computer simulation, for practical values of the operating BER.

Throughout this paper, the superscripts $T$ and $H$ represent the vector (matrix) transpose and conjugate transpose, respectively. $\Re\{x\}, \Im\{x\}$ and $\mathbb{E}[x]$ denote the real part, the imaginary part and the expected value of $x$, respectively, while $\|\mathbf{X}\|$ refers to the Frobenius norm of $\mathbf{X}$.

## II. Signal model

We consider a MIMO wireless communication system with 2 transmit antennas and $N_{\mathrm{r}}$ receive antennas. The propagation channels between both transmit antennas and each of the receive antennas are affected by flat fading with an arbitrary distribution. Transmission is organized in frames: in one frame, each transmit antenna sends $K_{\mathrm{p}}$ known pilot symbols and $K$ coded data symbols; the pilot symbols enable channel estimation at the receiver. Within one frame of $N_{\mathrm{fr}}=K+K_{\mathrm{p}}$ symbols, the channel is assumed to be constant (block fading). The $N_{\mathrm{r}} \times N_{\text {fr }}$ received signal matrix $\mathbf{R}_{\text {tot }}$ is given by

$$
\begin{equation*}
\mathbf{R}_{\mathrm{tot}}=\left[\mathbf{R}_{\mathrm{p}}, \mathbf{R}\right]=\mathbf{H}\left[\mathbf{A}_{\mathrm{p}}, \mathbf{A}\right]+\mathbf{W} \tag{1}
\end{equation*}
$$

where the $2 \times K_{\mathrm{p}}$ pilot matrix $\mathbf{A}_{\mathrm{p}}$ and the $2 \times K$ data matrix A consist of the pilot symbols and the coded data symbols, respectively, transmitted at each transmit antenna. The propagation channel is represented by the $N_{\mathrm{r}} \times 2$ complex random matrix $\mathbf{H}$. We consider an arbitrary joint pdf $p(\mathbf{H})$ of the (possibly correlated) $2 N_{\mathrm{r}}$ complex fading gains. The $N_{\mathrm{r}} \times N_{\mathrm{fr}}$ matrix $\mathbf{W}$ represents additive spatially and temporally white noise and consists of i.i.d. zero-mean circularly symmetric complex Gaussian (ZMCSCG) random variables with variance $N_{0}$.

Alamouti's space-time block code [1], that has been designed for two transmit antennas, transforms two information symbols $s_{1}(k)$ and $s_{2}(k)$ into a $2 \times 2$ coded symbol matrix $\mathbf{C}(k)$, given by

$$
\mathbf{C}(k)=\left[\begin{array}{cc}
s_{1}(k) & -s_{2}^{*}(k)  \tag{2}\\
s_{2}(k) & s_{1}^{*}(k)
\end{array}\right]
$$

Hence, assuming that an even number $K$ of information symbols is sent, the transmitted data symbol matrix is given by $\mathbf{A}=\sqrt{E_{\mathrm{s}}}[\mathbf{C}(1), \ldots, \mathbf{C}(K / 2)]$. Considering a normalized information symbol constellation $\left(\mathbb{E}\left[\left|s_{i}\right|^{2}\right]=1\right.$ ), it follows that the average energy of the transmitted coded symbols is given by $E_{\mathrm{s}}$ :

$$
\begin{equation*}
\frac{1}{2 K} \mathbb{E}\left[\|\mathbf{A}\|^{2}\right]=E_{\mathrm{s}} \tag{3}
\end{equation*}
$$

Similarly, the average energy of the pilot symbols is $E_{\mathrm{p}}$.

## III. Pilot-based channel estimation

The receiver estimates the channel matrix $\mathbf{H}$ using $\mathbf{R}_{\mathrm{p}}$ and the known pilot matrix $\mathbf{A}_{p}$. We assume that the rows of $\mathbf{A}_{p}$ are orthogonal, i.e. $\mathbf{A}_{\mathrm{p}} \mathbf{A}_{\mathrm{p}}^{H}=K_{\mathrm{p}} E_{\mathrm{p}} \mathbf{I}_{\mathrm{N}_{\mathrm{t}}}$, and consider linear channel estimates of the form

$$
\begin{equation*}
\hat{\mathbf{H}}=\frac{\alpha}{K_{\mathrm{p}} E_{\mathrm{p}}} \mathbf{R}_{\mathrm{p}} \mathbf{A}_{\mathrm{p}}^{H} \tag{4}
\end{equation*}
$$

with $\alpha \in \mathbb{R}$, such that $\hat{\mathbf{H}}$ can be decomposed into the sum of two statistically independent contributions:

$$
\begin{equation*}
\hat{\mathbf{H}}=\alpha \mathbf{H}+\mathbf{N} \tag{5}
\end{equation*}
$$

where the entries of $\mathbf{N}=\left(\alpha /\left(K_{\mathrm{p}} E_{\mathrm{p}}\right)\right) \mathbf{W}_{\mathrm{p}} \mathbf{A}_{\mathrm{p}}^{H}$ are ZMCSCG random variables; the real and imaginary parts of the entries of $\mathbf{N}$ have a variance $\sigma_{\mathbf{n}}^{2}=\alpha^{2} N_{0} /\left(2 K_{\mathrm{p}} E_{\mathrm{p}}\right)$. Hence, when conditioned on $\mathbf{H}$, the channel estimate $\hat{\mathbf{H}}$ is a complex Gaussian random matrix with mean $\alpha \mathbf{H}$ and diagonal covariance matrix with diagonal elements $2 \sigma_{\mathbf{n}}^{2}$. It is readily verified that both least-squares and linear MMSE estimation satisfy (4) with $\alpha=1$ and $\alpha=K_{\mathrm{p}} E_{\mathrm{p}} /\left(K_{\mathrm{p}} E_{\mathrm{p}}+N_{0}\right)$, respectively [16].

Allocating a large total energy $K_{\mathrm{p}} E_{\mathrm{p}}$ to pilot symbols yields an accurate channel estimate, but on the other hand gives rise to a reduction of the symbol energy $E_{\mathrm{s}}$. Denoting by $E_{\mathrm{b}}$ the energy per information bit and writing $E_{\mathrm{p}}=\gamma E_{\mathrm{s}}$, we have

$$
\begin{equation*}
E_{\mathrm{s}}=\frac{K}{K+\gamma K_{\mathrm{p}}} \rho \log _{2}(M) E_{\mathrm{b}} \tag{6}
\end{equation*}
$$

where $\rho=1 / 2$ is the ratio of the number of information symbols to the total number of symbols in Alamouti's code matrix, and M denotes the number of constellation points. Hence, $E_{\mathrm{s}}$ decreases with increasing $K_{\mathrm{p}}$.

## IV. ML Detection

When the channel state information (CSI) is available at the receiver, ML detection is known to be the optimal detection algorithm for the transmitted data:

$$
\begin{equation*}
\hat{\mathbf{A}}=\arg \min _{\tilde{\mathbf{A}}}\|\mathbf{R}-\mathbf{H} \tilde{\mathbf{A}}\|^{2} \tag{7}
\end{equation*}
$$

Denoting by $\left[\mathbf{r}_{1} \mathbf{r}_{2}\right]$ the $N_{r} \times 2$ observation matrix corresponding to a coded symbol matrix $\mathbf{C}$ (we omit the time index for notational convenience) and writing $\mathbf{H}=\left[\mathbf{h}_{1} \mathbf{h}_{2}\right]$, the detection algorithm for the information symbols $s_{1}$ and $s_{2}$ reduces to symbol-by-symbol detection:

$$
\begin{equation*}
\hat{s}_{i}=\arg \min _{s_{i}}\left|u_{i}-s_{i}\right|, i \in\{1,2\}, \tag{8}
\end{equation*}
$$

with

$$
\begin{align*}
& u_{1}=\frac{\mathbf{h}_{1}^{H} \mathbf{r}_{1}+\mathbf{h}_{2}^{T} \mathbf{r}_{2}^{*}}{\sqrt{E_{\mathrm{s}}}\|\mathbf{H}\|^{2}}  \tag{9}\\
& u_{2}=\frac{\mathbf{h}_{2}^{H} \mathbf{r}_{1}-\mathbf{h}_{1}^{T} \mathbf{r}_{2}^{*}}{\sqrt{E_{\mathrm{s}}}\|\mathbf{H}\|^{2}} \tag{10}
\end{align*}
$$

Since the CSI is not known by the receiver, we assume a mismatched receiver which uses the estimated channel matrix $\hat{\mathbf{H}}$ in the same way an ML receiver with PCK would use the
actual channel matrix $\mathbf{H}$. In this way, the decision variables become

$$
\begin{align*}
& u_{1}=\frac{\hat{\mathbf{h}}_{1}^{H} \mathbf{r}_{1}+\hat{\mathbf{h}}_{2}^{T} \mathbf{r}_{2}^{*}}{\sqrt{E_{\mathrm{s}}}\|\hat{\mathbf{H}}\|^{2}}  \tag{11}\\
& u_{2}=\frac{\hat{\mathbf{h}}_{2}^{H} \mathbf{r}_{1}-\hat{\mathbf{h}}_{1}^{T} \mathbf{r}_{2}^{*}}{\sqrt{E_{\mathrm{s}}}\|\hat{\mathbf{H}}\|^{2}} \tag{12}
\end{align*}
$$

In this contribution, we consider square $M$-QAM transmission with Gray mapping, which is equivalent to $\sqrt{M}-\mathrm{PAM}$ transmission for both the in-phase and quadrature information bits. The in-phase (quadrature) bits corresponding to the transmitted information symbol $s_{i}, i \in\{1,2\}$, are detected correctly when the real (imaginary) part of the decision variable $u_{i}$ is located inside the projection of the decision area of $s_{i}$ on the real (imaginary) axis. Hence, we can compute the BER of Alamouti's code as the average of the BERs for the in-phase and the quadrature bits of the symbols $s_{1}$ and $s_{2}$.

## V. Bit error rate analysis

Focusing our attention to the in-phase information bits related to the transmitted symbol $s_{1}$, the BER is obtained by averaging the conditional BER (conditioned on the channel $\mathbf{H}$ and the channel estimate $\hat{\mathbf{H}}$ ):

$$
\begin{equation*}
\mathrm{BER}_{1, \mathrm{R}}=\mathbb{E}_{\mathbf{H}, \hat{\mathbf{H}}}\left[\operatorname{BER}_{1, \mathrm{R}}(\mathbf{H}, \hat{\mathbf{H}})\right], \tag{13}
\end{equation*}
$$

with

$$
\begin{align*}
\operatorname{BER}_{1, \mathrm{R}}(\mathbf{H}, \hat{\mathbf{H}}) & = \\
\frac{1}{M^{2}} & \sum_{s_{1} \in \Psi} \sum_{s_{2} \in \Psi} \operatorname{BER}_{1, \mathrm{R}}\left(s_{1}, s_{2}, \mathbf{H}, \hat{\mathbf{H}}\right), \tag{14}
\end{align*}
$$

where $s_{1}$ and $s_{2}$ are the symbols actually transmitted, and $\Psi$ denotes the normalized $M$-QAM constellation. An inphase decision error occurs when the real part $u_{1, \mathrm{R}}$ of the decision variable $u_{1}$ is inside the projection (on the real axis) of the decision area of a QAM symbol $b$ for which the in-phase component $b_{\mathrm{R}}$ is different from the in-phase component $s_{1, \mathrm{R}}$ of the transmitted symbol $s_{1}$; this projection will be referred to as the decision region of $b_{\mathrm{R}}$. In this way, $\mathrm{BER}_{1, \mathrm{R}}\left(s_{1}, s_{2}, \mathbf{H}, \hat{\mathbf{H}}\right)$ is given by

$$
\begin{align*}
& \mathrm{BER}_{1, \mathrm{R}}\left(s_{1}, s_{2}, \mathbf{H}, \hat{\mathbf{H}}\right)= \\
& \quad \sum_{b_{\mathrm{R}} \in \Psi_{\mathrm{R}}} \frac{N\left(s_{1, \mathrm{R}}, b_{\mathrm{R}}\right)}{\frac{1}{2} \log _{2} M} P\left(s_{1}, s_{2}, b_{\mathrm{R}}, \mathbf{H}, \hat{\mathbf{H}}\right), \tag{15}
\end{align*}
$$

Here, $\Psi_{R}$ is the set consisting of the real parts of the constellation points, $N\left(s_{1, \mathrm{R}}, b_{\mathrm{R}}\right)$ represents the Hamming distance between the in-phase bits of the transmitted symbol $s_{1}$ and the in-phase bits of the detected symbol $b$, and $P\left(s_{1}, s_{2}, b_{\mathrm{R}}, \mathbf{H}, \hat{\mathbf{H}}\right)$ is the probability that the real part of the decision variable $u_{1}$ is located inside the decision area of $b_{\mathrm{R}}$ (when the transmitted symbols $s_{1}$ and $s_{2}$, the channel $\mathbf{H}$ and the channel estimate $\hat{\mathbf{H}}$ are known):

$$
\begin{equation*}
P\left(s_{1}, s_{2}, b_{\mathrm{R}}, \mathbf{H}, \hat{\mathbf{H}}\right)=\operatorname{Pr}\left[\hat{s}_{1, \mathrm{R}}=b_{\mathrm{R}} \mid s_{1}, s_{2}, \mathbf{H}, \hat{\mathbf{H}}\right] . \tag{16}
\end{equation*}
$$

Expanding the real part $u_{1, \mathrm{R}}$ of the decision variable (11) yields:

$$
\begin{equation*}
u_{1, \mathrm{R}}=u_{1, \mathrm{R}}^{\prime}+\frac{\Re\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{w}_{1}+\hat{\mathbf{h}}_{2}^{T} \mathbf{w}_{2}^{*}\right\}}{\sqrt{E_{\mathrm{s}}}\|\hat{\mathbf{H}}\|^{2}} \tag{17}
\end{equation*}
$$

with

$$
\begin{align*}
& u_{1, \mathrm{R}}^{\prime}=s_{1, \mathrm{R}} \frac{\Re\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{h}_{1}+\hat{\mathbf{h}}_{2}^{T} \mathbf{h}_{2}^{*}\right\}}{\|\hat{\mathbf{H}}\|^{2}} \\
& -s_{1, \mathrm{I}} \frac{\Im\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{h}_{1}+\hat{\mathbf{h}}_{2}^{T} \mathbf{h}_{2}^{*}\right\}}{\|\hat{\mathbf{H}}\|^{2}}+s_{2, \mathrm{R}} \frac{\Re\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{h}_{2}-\hat{\mathbf{h}}_{2}^{T} \mathbf{h}_{1}^{*}\right\}}{\|\hat{\mathbf{H}}\|^{2}} \\
& -s_{2, \mathrm{I}} \frac{\Im\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{h}_{2}-\hat{\mathbf{h}}_{2}^{T} \mathbf{h}_{1}^{*}\right\}}{\|\hat{\mathbf{H}}\|^{2}} \tag{18}
\end{align*}
$$

In (18), the subscripts R and I refer to the real and imaginary parts of the transmitted symbols, respectively. The second term in (17) represents zero-mean Gaussian noise with variance $N_{0} /\left(2 E_{\mathrm{s}}\|\hat{\mathbf{H}}\|^{2}\right)$. The first term in (18) is the useful term, whereas the second term represents interference from the quadrature component $s_{1, \mathrm{I}}$. The third and fourth term represent interference from the in-phase and quadrature component of the transmitted symbol $s_{2}$, respectively. Note that $u_{1, \mathrm{R}}^{\prime}=s_{1, \mathrm{R}}$ when $\hat{\mathbf{H}}=\mathbf{H}$ (i.e. for PCK). Let $g_{1}\left(b_{\mathrm{R}}\right)$ and $g_{2}\left(b_{\mathrm{R}}\right)$ denote the boundaries of the projection of the decision area of $b$ on the real axis, with $g_{1}\left(b_{\mathrm{R}}\right)<g_{2}\left(b_{\mathrm{R}}\right)$; we set $g_{1}\left(b_{\mathrm{R}}\right)=-\infty$ $\left(g_{2}\left(b_{\mathrm{R}}\right)=\infty\right)$ if $b$ is a left (right) outer constellation point. In this way, (16) reduces to

$$
\begin{equation*}
P\left(s_{1}, s_{2}, b_{\mathrm{R}}, \mathbf{H}, \hat{\mathbf{H}}\right)=Q_{1}-Q_{2} \tag{19}
\end{equation*}
$$

where the quantities $Q_{k}, k \in\{1,2\}$, are given by

$$
\begin{equation*}
Q_{k}=Q\left(\sqrt{2 \frac{E_{\mathrm{s}}}{N_{0}}}\|\hat{\mathbf{H}}\|\left(g_{k}\left(b_{\mathrm{R}}\right)-u_{1, \mathrm{R}}^{\prime}\right)\right) \tag{20}
\end{equation*}
$$

In (20), $Q($.$) is the Gaussian \mathrm{Q}$-function, defined as

$$
\begin{equation*}
Q(x) \triangleq \frac{1}{\sqrt{2 \pi}} \int_{x}^{\infty} \exp \left(-\frac{x^{2}}{2}\right) \mathrm{d} x \tag{21}
\end{equation*}
$$

Hence, (13) can be expressed as

$$
\begin{equation*}
\mathrm{BER}_{1, \mathrm{R}}=\int \operatorname{BER}_{1, \mathrm{R}}(\mathbf{H}, \hat{\mathbf{H}}) p(\hat{\mathbf{H}} \mid \mathbf{H}) p(\mathbf{H}) \mathrm{d} \hat{\mathbf{H}} \mathrm{~d} \mathbf{H} \tag{22}
\end{equation*}
$$

where $p(\hat{\mathbf{H}} \mid \mathbf{H})$ is the joint Gausian pdf of $2 N_{\mathrm{r}}$ complex-valued random variables, with mean $\alpha \mathbf{H}$ and diagonal covariance matrix with diagonal elements $2 \sigma_{\mathbf{n}}^{2}$, and $p(\mathbf{H})$ represents the arbitrary fading distribution. Note that this general description also allows correlation between the components of $\mathbf{H}$. Although (22) is conceptually simple, it is not well suited for numerical evaluation. Indeed, taking into account that the components of $\mathbf{H}$ and $\hat{\mathbf{H}}$ are complex-valued, the evaluation of (22) requires taking the expectation over $8 N_{\mathrm{r}}$ real-valued variables; hence, the associated computing time increases dramatically with the number $N_{\mathrm{r}}$ of receive antennas.

In order to avoid the computational complexity corresponding to the evaluation of (22), we will manipulate (22) into an
expectation over only 6 variables. To this end, we introduce the following real-valued vectors:

$$
\left\{\begin{array}{l}
\hat{\mathbf{h}}^{\prime}=\left[\hat{\mathbf{h}}_{1, \mathrm{R}}^{T}, \hat{\mathbf{h}}_{1, \mathrm{I}}^{T}, \hat{\mathbf{h}}_{2, \mathrm{R}}^{T}, \hat{\mathbf{h}}_{2, \mathrm{I}}^{T}\right]^{T}  \tag{23}\\
\mathbf{h}_{1}^{\prime}=\left[\mathbf{h}_{1, \mathrm{R}}^{T}, \mathbf{h}_{1, \mathrm{I}}^{T}, \mathbf{h}_{2, \mathrm{R}}^{T}, \mathbf{h}_{2, \mathrm{I}}^{T}\right]^{T} \\
\mathbf{h}_{2}^{\prime}=\left[\mathbf{h}_{1, \mathrm{I}}^{T},-\mathbf{h}_{1, \mathrm{R}}^{T},-\mathbf{h}_{2, \mathrm{I}}^{T}, \mathbf{h}_{2, \mathrm{R}}^{T}\right]^{T} \\
\mathbf{h}_{3}^{\prime}=\left[\mathbf{h}_{2, \mathrm{R}}^{T}, \mathbf{h}_{2, \mathrm{I}}^{T},-\mathbf{h}_{1, \mathrm{R}}^{T},-\mathbf{h}_{1, \mathrm{I}}^{T}\right]^{T} \\
\mathbf{h}_{4}^{\prime}=\left[\mathbf{h}_{2, \mathrm{I}}^{T},-\mathbf{h}_{2, \mathrm{R}}^{T}, \mathbf{h}_{1, \mathrm{I}}^{T},-\mathbf{h}_{1, \mathrm{R}}^{T}\right]^{T}
\end{array},\right.
$$

with $\hat{\mathbf{h}}_{i, \mathrm{R}}+j \hat{\mathbf{h}}_{i, \mathrm{I}}$ and $\mathbf{h}_{i, \mathrm{R}}+j \mathbf{h}_{i, \mathrm{I}}, i \in\{1,2\}$, denoting the $i$-th column of $\hat{\mathbf{H}}$ and $\mathbf{H}$, respectively. It is easily verified that $\mathbf{h}_{1}^{\prime}, \mathbf{h}_{2}^{\prime}, \mathbf{h}_{3}^{\prime}$ and $\mathbf{h}_{4}^{\prime}$ are orthogonal. Let us consider an orthonormal coordinate system defined by the unit vectors $\left\{\mathbf{e}_{i}, i=1, \cdots, 4 N_{\mathrm{r}}\right\}$, with $\mathbf{e}_{1}, \mathbf{e}_{2}, \mathbf{e}_{3}$ and $\mathbf{e}_{4}$ directed along $\mathbf{h}_{1}^{\prime}, \mathbf{h}_{2}^{\prime}, \mathbf{h}_{3}^{\prime}$ and $\mathbf{h}_{4}^{\prime}$, respectively. The projections of $\hat{\mathbf{h}}^{\prime}$ on $\mathbf{e}_{i}$ are denoted $x_{i}$, with $x_{i}=\hat{\mathbf{h}}^{T} \mathbf{e}_{\mathbf{i}}$, yielding $\|\hat{\mathbf{H}}\|^{2}=\left|\hat{\mathbf{h}}^{\prime}\right|^{2}=$ $x_{1}^{2}+x_{2}^{2}+x_{3}^{2}+x_{4}^{2}+z^{2}$, where

$$
\begin{equation*}
z^{2}=\sum_{i=5}^{4 N_{\mathrm{r}}} x_{i}^{2} \tag{24}
\end{equation*}
$$

Because of the specific choice of $\mathbf{e}_{i}$ for $1 \leq i \leq 4$, we have

$$
\begin{align*}
& x_{1}=\frac{\hat{\mathbf{h}}^{T} \mathbf{h}_{\mathbf{1}}^{\prime}}{\left|\mathbf{h}_{\mathbf{1}}^{\prime}\right|}=\frac{\Re\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{h}_{1}+\hat{\mathbf{h}}_{2}^{T} \mathbf{h}_{2}^{*}\right\}}{\|\mathbf{H}\|},  \tag{25}\\
& x_{2}=\frac{\hat{\mathbf{h}}^{T} \mathbf{h}_{\mathbf{2}}^{\prime}}{\left|\mathbf{h}_{\mathbf{2}}^{\prime}\right|}=\frac{\Im\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{h}_{1}+\hat{\mathbf{h}}_{2}^{T} \mathbf{h}_{2}^{*}\right\}}{\|\mathbf{H}\|},  \tag{26}\\
& x_{3}=\frac{\hat{\mathbf{h}}^{T} \mathbf{h}_{\mathbf{3}}^{\prime}}{\left|\mathbf{h}_{\mathbf{3}}^{\prime}\right|}=\frac{\Re\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{h}_{2}-\hat{\mathbf{h}}_{2}^{T} \mathbf{h}_{1}^{*}\right\}}{\|\mathbf{H}\|},  \tag{27}\\
& x_{4}=\frac{\hat{\mathbf{h}}^{T} \mathbf{h}_{\mathbf{4}}^{\prime}}{\left|\mathbf{h}_{\mathbf{4}}^{\prime}\right|}=\frac{\Im\left\{\hat{\mathbf{h}}_{1}^{H} \mathbf{h}_{2}-\hat{\mathbf{h}}_{2}^{T} \mathbf{h}_{1}^{*}\right\}}{\|\mathbf{H}\|} . \tag{28}
\end{align*}
$$

Let us introduce the vector $\mathbf{x}=\left(x_{1}, x_{2}, x_{3}, x_{4}\right)$. From (18) and (20) it follows that $\operatorname{BER}_{1, \mathrm{R}}(\mathbf{H}, \hat{\mathbf{H}})$ is a function of only 6 random variables, $x_{1}, x_{2}, x_{3}, x_{4}, z$ and $\|\mathbf{H}\|$, which we denote $\mathrm{BER}_{1, \mathrm{R}}(\mathbf{x}, z,\|\mathbf{H}\|)$. Taking (5) into account, it can be shown that $x_{1}, x_{2}, x_{3}, x_{4}$ and $z$ (when conditioned on $\|\mathbf{H}\|$ ) are independent variables which satisfy the following properties:

- $x_{1}$ is a Gaussian variable with mean $\alpha\|\mathbf{H}\|$ and variance $\sigma_{\mathrm{n}}^{2}$.
- $x_{2}, x_{3}$ and $x_{4}$ are zero-mean Gaussian variables with variance $\sigma_{\mathbf{n}}^{2}$.
- $z / \sigma_{\mathbf{n}}$ is distributed according to the chi-distribution with $4 N_{\mathrm{r}}-4$ degrees of freedom [3].
Taking the above properties into account, $\mathrm{BER}_{1, \mathrm{R}}$ is obtained as

$$
\begin{align*}
& \operatorname{BER}_{1, \mathrm{R}}=\int \operatorname{BER}_{1, \mathrm{R}}(\mathbf{x}, z, u) \\
& \qquad p(\mathbf{x}, z \mid\|\mathbf{H}\|=u) p(u) \mathrm{d} \mathbf{x} \mathrm{~d} z \mathrm{~d} u \tag{29}
\end{align*}
$$

where $p(u)$ is the pdf of the Frobenius norm $\|\mathbf{H}\|$ of the channel matrix $\mathbf{H}$. It is important to note that instead of the joint distribution $p(\mathbf{H})$ of the $2 N_{\mathbf{r}}$ complex-valued fading
gains, we need only the distribution of the Frobenius norm $\|\mathbf{H}\|$. Owing to the rotational symmetry of the $M$-QAM constellation and taking into account that the symbol vectors are equally likely, it can be verified that the BERs related to the in-phase and quadrature bits of $s_{1}$ and $s_{2}$ are equal. Hence, the $B E R$ resulting from Alamouti's code equals $\mathrm{BER}_{1, \mathrm{R}}$, and the BER evaluation involves an expectation over only 6 random variables. The expectation (29) can be evaluated numerically by approximating the 6 -fold integral by a 6 -fold sum, running over discretized versions of the continuous variables $x_{1}, x_{2}$, $x_{3}, x_{4}, z$ and $u$.

The number of random variables to be considered in the expectation (29) is reduced in the following cases:

1) In the case of single-input single-output (SISO) transmission $\left(N_{\mathrm{r}}=1\right)$, the dimension of $\hat{\mathbf{h}}^{\prime}$ is 4 , so the only projections are $x_{1}, x_{2}, x_{3}$ and $x_{4}$. Hence the BER computation involves averaging over only 5 variables, $x_{1}, x_{2}, x_{3}, x_{4}$ and $\|\mathbf{H}\|$.
2) For PAM constellations, we have $s_{1, \mathrm{I}}=s_{2, \mathrm{I}}=0$, so that $u_{1, \mathrm{R}}^{\prime}$ from (18) does not depend on $x_{2}$ and $x_{4}$. Hence, the BER (which equals the in-phase BER) is obtained by averaging over only 4 variables, $x_{1}, x_{3}, z^{\prime}$ and $\|\mathbf{H}\|$, where $\|\hat{\mathbf{H}}\|^{2}=x_{1}^{2}+x_{3}^{2}+z^{\prime 2}$ and $z^{\prime} / \sigma_{\mathbf{n}}$ has a chidistribution with $4 N_{\mathrm{r}}-2$ degrees of freedom.

Till now we have considered complex-valued channel gains. In some applications (e.g., ultra wideband communication), channel models with real-valued gains are more appropriate (and the transmitted symbols are real-valued too). The BER result in this case is nearly the same as for PAM transmission over complex-valued channels (averaging over 4 variables, $x_{1}$, $x_{3}, z^{\prime}$ and $\|\mathbf{H}\|$ ), the only difference being that on real-valued channels the variable $z^{\prime} / \sigma_{\mathbf{n}}$ has $2 N_{\mathrm{r}}-2$ (instead of $4 N_{\mathrm{r}}-2$ ) degrees of freedom. For specific cases, a further reduction of the number of random variables can be obtained. According to [6], only a twofold integral must be computed in the case of $M$-PSK constellations and arbitrary fading.

## VI. Numerical results

To obtain our numerical results, we assume MMSE channel estimation with $E_{\mathrm{p}}=E_{\mathrm{s}}$ and $2 N_{\mathrm{r}}$ independent and identically distributed (i.i.d.) Nakagami- $m$ fading channels, with parameters $m$ and $\Omega$ [14].

Fig. 1 shows the exact BER curves for Alamouti's code and 16-QAM transmission over a complex-valued MIMO channel, for both the mismatched receiver and the PCK receiver, and for several values of $m$ and $N_{\mathrm{r}}$. It is easily derived that the Frobenius norm $\|\mathbf{H}\|$ of the MIMO channel follows a Nakagami- $m$ distribution with parameters $2 N_{\mathrm{r}} m$ and $2 N_{\mathrm{r}} \Omega$. The curves corresponding to the mismatched receiver represent the numerically computed expectation over 6 variables resulting from (29), whereas the BER of the PCK receiver involves the expectation over $\|\mathbf{H}\|$ only. Also shown in the figure are straightforward computer simulation results for the mismatched receiver that confirm the result obtained from (but require considerably more computing time then) numerical averaging.


Fig. 1. Complex-valued Nakagami-m fading channel, 16-QAM

## VII. CONCLUSIONS AND REMARKS

In this contribution, we investigated the effect of imperfect channel estimation on the BER performance of Alamouti's code. The transmitted symbols belong to a PAM or QAM constellation. We assumed pilot symbol assisted channel estimation and $2 N_{\mathrm{r}}$ propagation channels affected by flat block fading with arbitrary pdf. The main conclusions are the following:

- A conceptually simple BER expression can be obtained in the form of an expectation over $8 N_{\mathrm{r}}$ random variables. However, its numerical evaluation requires a computing time that increases dramatically with $N_{\mathrm{r}}$.
- We have reduced the BER expression to a form that contains an expectation over only 6 variables, irrespective of the value of $N_{\mathrm{r}}$. Rather than the joint distribution of all $2 N_{\mathrm{r}}$ complex-valued channel gains, this expression depends on the distribution of only the Frobenius norm of the channel matrix. We have pointed out that for some cases (SISO, real-valued channel, PAM constellation, and combinations thereof) the number of variables involved in the expectation can be further reduced.
- Evaluation of the reduced BER expression generally requires a ( 6 -fold or less) numerical integration. This BER expression can be used when the pdf of the Frobenius norm of the channel matrix is available in closed form or has been determined experimentally (histogram). Comparing the computing times resulting from the numerical averaging and from straightforward simulation, it turns out that the numerical averaging is to be preferred for BER values of practical interest.
The reviewers brought to our attention that OSTBCs can be represented by an equivalent SISO channel, which allows to exploit SISO channel results for studying OSTBCs [13], [12], [7]. This avenue will be explored in further research.


## Acknowledgments

This work was supported by the European Commission in the framework of the FP7 Network of Excellence in Wireless

COMmunications NEWCOM++ (contract no. 216715). The first author also gratefully acknowledges the support from the Fund for Scientific Research in Flanders (FWO-Vlaanderen).

## References

[1] S.M. Alamouti. A simple transmit diversity technique for wireless communications. IEEE J. Select. Areas Commun., 16:1459-1478, Oct 1998.
[2] M. Brehler and M.K. Varanasi. Asymptotic error probability analysis of quadratic receivers in rayleigh-fading channels with applications to a unified analysis of coherent and noncoherent space-time receivers. IEEE Trans. Inform. Theory, 47(6):2383-2399, Sept. 2001.
[3] M. Evans, N. Hastings, and B. Peacock. Statistical Distributions 3rd ed. Wiley, New York, 2000.
[4] G. Femenias. BER performance of linear STBC from orthogonal designs over MIMO correlated Nakagami- $m$ fading channels. IEEE Trans. Veh. Technol., 53(2):307-317, March 2004.
[5] P. Garg, R.K. Mallik, and H.M. Gupta. Performance analysis of spacetime coding with imperfect channel estimation. IEEE Trans. Wireless Commun., 4(1):257-265, Jan. 2005.
[6] P. Garg, R.K. Mallik, and H.M. Gupta. Exact error performance of square orthogonal space-time block coding with channel estimation. IEEE Trans. Commun., 54(3):430-437, March 2006.
[7] W.M. Gifford, M.Z. Win, and M. Chiani. Diversity with practical channel estimation. IEEE Trans. Wireless Commun., 4(4):1935-1947, July 2005.
[8] L. Jacobs and M. Moeneclaey. Effect of MMSE channel estimation on BER performance of orthogonal space-time block codes in Rayleigh fading channels. Accepted for publication in IEEE Trans. Commun.
[9] I.-M. Kim. Exact BER analysis of OSTBCs in spatially correlated MIMO channels. IEEE Trans. Commun., 54(8):1365-1373, Aug. 2006.
[10] I.-M. Kim and V. Tarokh. Variable-rate space-time block codes in Mary PSK systems. IEEE J. Select. Areas Commun., 21(3):362-373, April 2003.
[11] S.-H. Kim, I.-S. Kang, and J.-S. No. Exact bit error probability of orthogonal space-time block codes with QAM. In Proc. IEEE Int. Symp. on Information Theory (ISIT), page 63, Yokohama, Japan, 29 June-4 July 2003.
[12] A. Maaref and S. Aissa. Performance analysis of orthogonal space-time block codes in spatially correlated MIMO Nakagami fading channels. IEEE Trans. Wireless Commun., 5(4):807-817, April 2006.
[13] A. Maaref and S. Aissa. Capacity of space-time block codes in mimo Rayleigh fading channels with adaptive transmission and estimation errors. IEEE Trans. Wireless Commun., 4(5):2568-2578, Sept. 2006.
[14] M. Nakagami. The $m$-distribution-A general formula for intensity distribution of rapid fading. In Statistical Methods in Radio Wave Propagation, pages 3-36. New York: Pergamon, 1960.
[15] C. Shan, Kam P.Y., and A. Nallanathan. Theoretical performance of space-time block coded system with channel estimation. In Proc. IEEE Global Telecommunications Conf. (Globecom), pages 3666-3670 Vol.6, Dallas, USA, 29 Nov.-3 Dec. 2004.
[16] G. Taricco and E. Biglieri. Space-time decoding with imperfect channel estimation. IEEE Trans. Wireless Commun., 4(4):1874-1888, July 2005.
[17] V. Tarokh, H. Jafarkhani, and A.R. Calderbank. Space-time block codes from orthogonal designs. IEEE Trans. Inform. Theory, 45(5):1456-1467, July 1999.
[18] Liang X.-B. Orthogonal designs with maximal rates. IEEE Trans. Inform. Theory, 49(10):2468-2503, Oct. 2003.

