

# E016350 - Artificial Intelligence

## Lecture 15

### **Reasoning under Uncertainty & Bayesian ML** Temporal probability models

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# Overview

- Time and uncertainty
- Inference: filtering, prediction, smoothing
- Hidden Markov models
- Kalman filters (a brief mention)
- Dynamic Bayesian networks
- Particle filtering

[R&N], Chapter 14

This presentation is based on: S. Russel and P. Norvig: *Artificial Intelligence: A Modern Approach*, (Fourth Ed.), denoted as [R&N] and corresp. resources <http://aima.cs.berkeley.edu/>

# Time and uncertainty

The world changes; we need to track and predict it

E.g., diabetes management (dynamic) vs vehicle diagnosis (static)

Basic idea: keep track of state and evidence variables at each time step

$\mathbf{X}_t$  = set of unobservable state variables at time  $t$   
e.g., *BloodSugar<sub>t</sub>*, *StomachContents<sub>t</sub>*, etc.

$\mathbf{E}_t$  = set of observable evidence variables at time  $t$   
e.g., *MeasuredBloodSugar<sub>t</sub>*, *PulseRate<sub>t</sub>*, *FoodEaten<sub>t</sub>*

This assumes **discrete time**; step size depends on problem

Notation:  $\mathbf{X}_{a:b} = \mathbf{X}_a, \mathbf{X}_{a+1}, \dots, \mathbf{X}_{b-1}, \mathbf{X}_b$

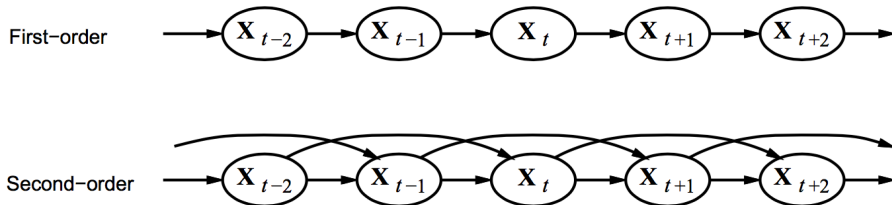
# Markov processes (Markov chains)

Construct a Bayes net from these variables: parents?

Markov assumption:  $\mathbf{X}_t$  depends on **bounded subset** of  $\mathbf{X}_{0:t-1}$

First-order Markov process:  $\mathbf{P}(\mathbf{X}_t | \mathbf{X}_{0:t-1}) = \mathbf{P}(\mathbf{X}_t | \mathbf{X}_{t-1})$

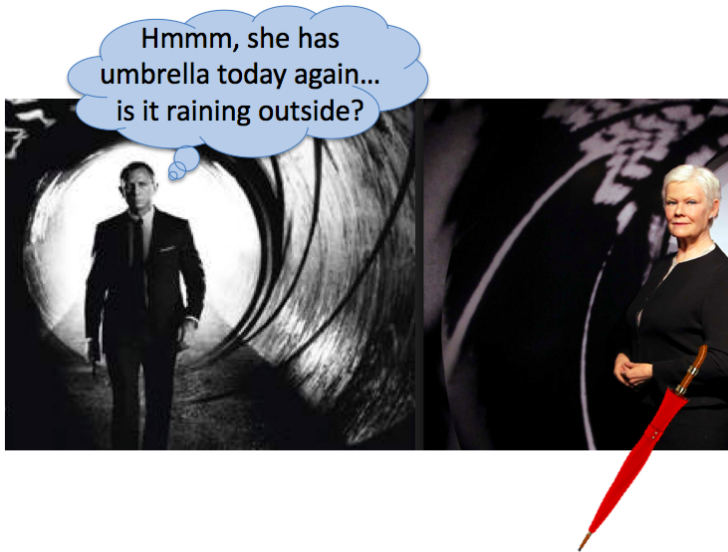
Second-order Markov process:  $\mathbf{P}(\mathbf{X}_t | \mathbf{X}_{0:t-1}) = \mathbf{P}(\mathbf{X}_t | \mathbf{X}_{t-2}, \mathbf{X}_{t-1})$



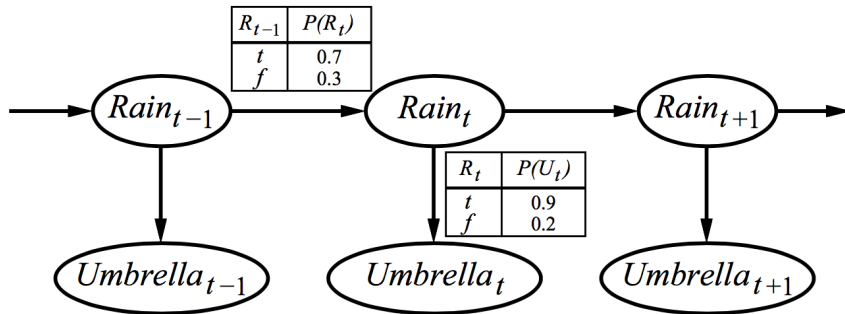
Sensor Markov assumption:  $\mathbf{P}(\mathbf{E}_t | \mathbf{X}_{0:t}, \mathbf{E}_{0:t-1}) = \mathbf{P}(\mathbf{E}_t | \mathbf{X}_t)$

Stationary process: transition model  $\mathbf{P}(\mathbf{X}_t | \mathbf{X}_{t-1})$  and  
sensor model  $\mathbf{P}(\mathbf{E}_t | \mathbf{X}_t)$  fixed for all  $t$

## Example: umbrella



## Example: umbrella



First-order Markov assumption not exactly true in real world!

Possible fixes:

1. **Increase order** of Markov process
2. **Augment state**, e.g., add  $Temp_t$ ,  $Pressure_t$

Another example: robot motion. Augment position and velocity with  $Battery_t$

# Inference tasks

Filtering:  $\mathbf{P}(\mathbf{X}_t | \mathbf{e}_{1:t})$

belief state – input to the decision process of a rational agent

Prediction:  $\mathbf{P}(\mathbf{X}_{t+k} | \mathbf{e}_{1:t})$  for  $k > 0$

evaluation of possible action sequences;  
like filtering without the evidence

Smoothing:  $\mathbf{P}(\mathbf{X}_k | \mathbf{e}_{1:t})$  for  $0 \leq k < t$

better estimate of past states, essential for learning

Most likely explanation:  $\arg \max_{\mathbf{x}_{1:t}} P(\mathbf{x}_{1:t} | \mathbf{e}_{1:t})$

speech recognition, decoding with a noisy channel

# Filtering

Aim: devise a **recursive** state estimation algorithm:

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = f(\mathbf{e}_{t+1}, \mathbf{P}(\mathbf{X}_t|\mathbf{e}_{1:t}))$$

$$\begin{aligned}\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) &= \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t}, \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1}, \mathbf{e}_{1:t}) \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1}) \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t})\end{aligned}$$



# Filtering

Aim: devise a **recursive** state estimation algorithm:

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = f(\mathbf{e}_{t+1}, \mathbf{P}(\mathbf{X}_t|\mathbf{e}_{1:t}))$$

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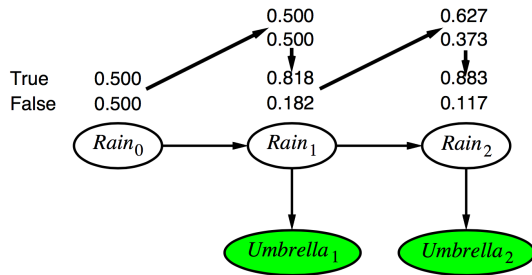
I.e., **prediction** + **estimation**. Prediction by summing out  $\mathbf{X}_t$ :

$$\begin{aligned}\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) &= \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1}) \sum_{\mathbf{x}_t} \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{x}_t, \mathbf{e}_{1:t}) P(\mathbf{x}_t|\mathbf{e}_{1:t}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1}) \sum_{\mathbf{x}_t} \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{x}_t) P(\mathbf{x}_t|\mathbf{e}_{1:t})\end{aligned}$$

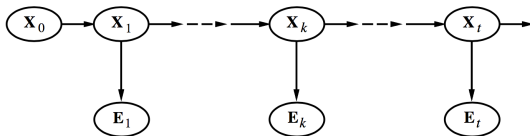
$\mathbf{f}_{1:t+1} = \alpha \text{FORWARD}(\mathbf{f}_{1:t}, \mathbf{e}_{t+1})$  where  $\mathbf{f}_{1:t} = \mathbf{P}(\mathbf{X}_t|\mathbf{e}_{1:t})$

Time and space **constant** (independent of  $t$ )

## Filtering example



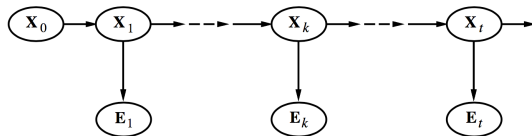
# Smoothing



Divide evidence  $\mathbf{e}_{1:t}$  into  $\mathbf{e}_{1:k}$ ,  $\mathbf{e}_{k+1:t}$ :

$$\begin{aligned} \mathbf{P}(\mathbf{X}_k | \mathbf{e}_{1:t}) &= \mathbf{P}(\mathbf{X}_k | \mathbf{e}_{1:k}, \mathbf{e}_{k+1:t}) = \alpha \mathbf{P}(\mathbf{X}_k | \mathbf{e}_{1:k}) \mathbf{P}(\mathbf{e}_{k+1:t} | \mathbf{X}_k, \mathbf{e}_{1:k}) \\ &= \alpha \mathbf{P}(\mathbf{X}_k | \mathbf{e}_{1:k}) \mathbf{P}(\mathbf{e}_{k+1:t} | \mathbf{X}_k) \\ &= \alpha \mathbf{f}_{1:k} \mathbf{b}_{k+1:t} \end{aligned}$$

# Smoothing



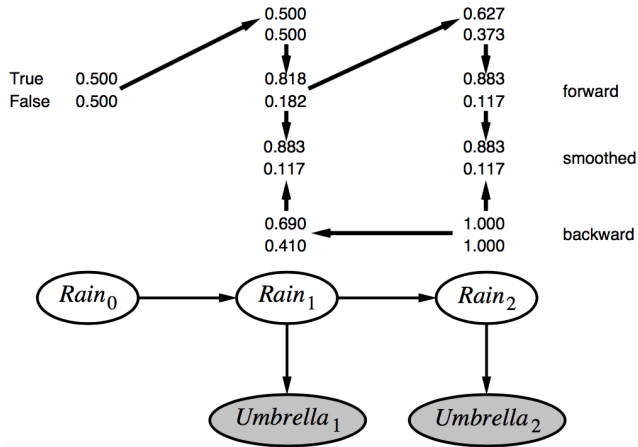
Divide evidence  $\mathbf{e}_{1:t}$  into  $\mathbf{e}_{1:k}$ ,  $\mathbf{e}_{k+1:t}$ :

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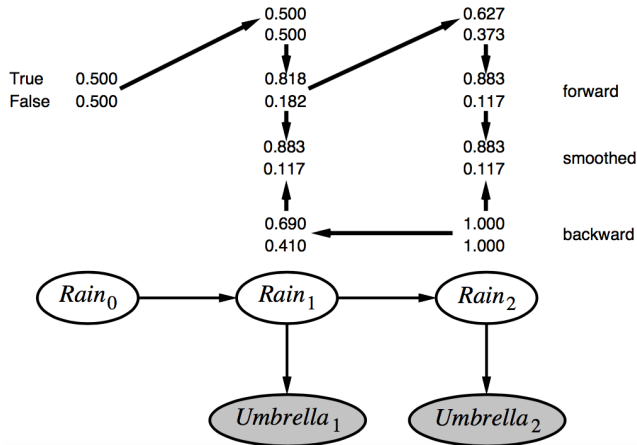
Backward message computed by a backwards recursion:

$$\begin{aligned}\mathbf{P}(\mathbf{e}_{k+1:t} | \mathbf{X}_k) &= \sum_{\mathbf{x}_{k+1}} \mathbf{P}(\mathbf{e}_{k+1:t} | \mathbf{X}_k, \mathbf{x}_{k+1}) \mathbf{P}(\mathbf{x}_{k+1} | \mathbf{X}_k) \\ &= \sum_{\mathbf{x}_{k+1}} P(\mathbf{e}_{k+1:t} | \mathbf{x}_{k+1}) \mathbf{P}(\mathbf{x}_{k+1} | \mathbf{X}_k) \\ &= \sum_{\mathbf{x}_{k+1}} P(\mathbf{e}_{k+1} | \mathbf{x}_{k+1}) P(\mathbf{e}_{k+2:t} | \mathbf{x}_{k+1}) \mathbf{P}(\mathbf{x}_{k+1} | \mathbf{X}_k)\end{aligned}$$

# Smoothing example



# Smoothing example



Forward-backward algorithm: cache forward messages along the way  
Time linear in  $t$  (polytree inference), space  $O(t|f|)$

## Most likely explanation

Most likely sequence  $\neq$  sequence of most likely states!!!!

Most likely path to each  $\mathbf{x}_{t+1}$

= most likely path to **some**  $\mathbf{x}_t$  plus one more step

$$\begin{aligned} & \max_{\mathbf{x}_1 \dots \mathbf{x}_t} \mathbf{P}(\mathbf{x}_1, \dots, \mathbf{x}_t, \mathbf{X}_{t+1} | \mathbf{e}_{1:t+1}) \\ &= \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{X}_{t+1}) \max_{\mathbf{x}_t} \left( \mathbf{P}(\mathbf{X}_{t+1} | \mathbf{x}_t) \max_{\mathbf{x}_1 \dots \mathbf{x}_{t-1}} P(\mathbf{x}_1, \dots, \mathbf{x}_{t-1}, \mathbf{x}_t | \mathbf{e}_{1:t}) \right) \end{aligned}$$

Identical to filtering, except  $\mathbf{f}_{1:t}$  replaced by

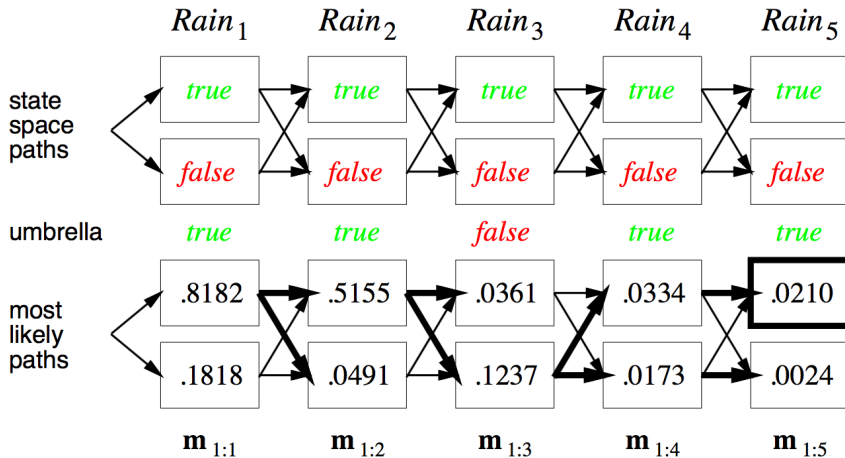
$$\mathbf{m}_{1:t} = \max_{\mathbf{x}_1 \dots \mathbf{x}_{t-1}} \mathbf{P}(\mathbf{x}_1, \dots, \mathbf{x}_{t-1}, \mathbf{X}_t | \mathbf{e}_{1:t}),$$

I.e.,  $\mathbf{m}_{1:t}(i)$  gives the probability of the most likely path to state  $i$ .

Update has sum replaced by max, giving the **Viterbi algorithm**:

$$\mathbf{m}_{1:t+1} = \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{X}_{t+1}) \max_{\mathbf{x}_t} (\mathbf{P}(\mathbf{X}_{t+1} | \mathbf{x}_t) \mathbf{m}_{1:t})$$

# Viterbi example





# Hidden Markov models

$\mathbf{X}_t$  is a single, discrete variable (usually  $\mathbf{E}_t$  is too)

Domain of  $\mathbf{X}_t$  is  $\{1, \dots, S\}$

Transition matrix  $\mathbf{T}_{ij} = P(X_t = j | X_{t-1} = i)$ ,

e.g., for the umbrella world  $\mathbf{T} = \begin{pmatrix} 0.7 & 0.3 \\ 0.3 & 0.7 \end{pmatrix}$

Sensor matrix  $\mathbf{O}_t$  for each time step, diagonal elements  $P(e_t | X_t = i)$

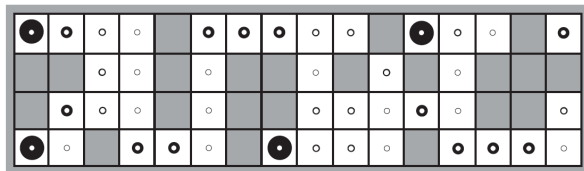
e.g., for the umbrella world with  $U_1 = \text{true}$ ,  $\mathbf{O}_1 = \begin{pmatrix} 0.9 & 0 \\ 0 & 0.2 \end{pmatrix}$

Forward and backward messages as column vectors:

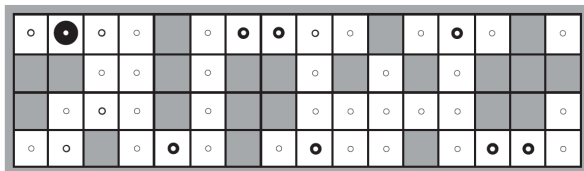
$$\begin{aligned} \mathbf{f}_{1:t+1} &= \alpha \mathbf{O}_{t+1} \mathbf{T}^\top \mathbf{f}_{1:t} \\ \mathbf{b}_{k+1:t} &= \mathbf{T} \mathbf{O}_{k+1} \mathbf{b}_{k+2:t} \end{aligned}$$

Forward-backward algorithm needs time  $O(S^2t)$  and space  $O(St)$

## Example: robot localization



(a) Posterior distribution over robot location after  $E_1 = \text{NSW}$



(b) Posterior distribution over robot location after  $E_1 = \text{NSW}, E_2 = \text{NS}$

Transition model:

$$P(X_{t+1} = j | X_t = i) = \mathbf{T}_{i,j} = \frac{1}{N(i)}$$

if  $j \in \text{NEIGHBORS}(i)$ , else 0

Sensor model:

$$P(E_t = e_t | X_t = i) = \mathbf{O}_{t,i,i} \\ = (1 - \epsilon)^{4-d_{it}} \epsilon^{d_{it}}$$

$d_{it}$  is the discrepancy (the number of bits that are different between the true values for square  $i$  and the actual reading  $e_t$ );  $\epsilon$  – sensor error rate

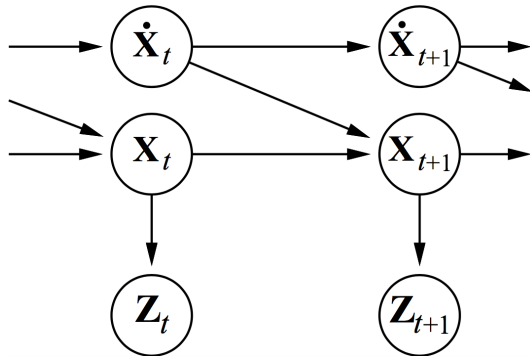
Posterior distribution  $P(X_t = i | e_t)$  over robot location: (a) one observation  $E_1 = \text{NSW}$ ; (b) after a second observation  $E_2 = \text{NS}$ . The size of each disk corresponds to the probability that the robot is at that location.  $\epsilon = 0.2$

## Kalman filters

Modelling systems described by a set of continuous variables,

e.g., tracking a bird flying— $\mathbf{X}_t = X, Y, Z, \dot{X}, \dot{Y}, \dot{Z}$ .

Airplanes, robots, ecosystems, economies, chemical plants, planets, ...



Gaussian prior, linear Gaussian transition model and sensor model

## Updating Gaussian distributions

- 1) If the current distribution  $\mathbf{P}(\mathbf{X}_t|\mathbf{e}_{1:t})$  is Gaussian, and the transition model  $\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{x}_t)$  is linear Gaussian, then prediction

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t}) = \int_{\mathbf{x}_t} \mathbf{P}(\mathbf{X}_{t+1}|\mathbf{x}_t)P(\mathbf{x}_t|\mathbf{e}_{1:t}) d\mathbf{x}_t$$

is also Gaussian.

- 2) If the prediction  $\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t})$  is Gaussian, and the sensor model  $\mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_t)$  is linear Gaussian, then the updated distribution after conditioning on new evidence

$$\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1}|\mathbf{X}_{t+1})\mathbf{P}(\mathbf{X}_{t+1}|\mathbf{e}_{1:t})$$

is also a Gaussian distribution

Hence  $\mathbf{P}(\mathbf{X}_t|\mathbf{e}_{1:t})$  is multivariate Gaussian  $N(\boldsymbol{\mu}_t, \boldsymbol{\Sigma}_t)$  for all  $t$

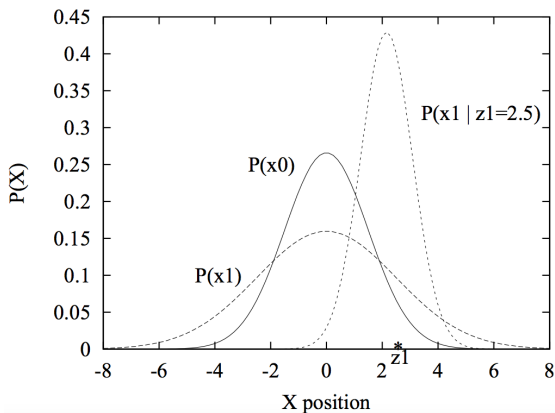
General (nonlinear, non-Gaussian) process: description of posterior grows **unboundedly** as  $t \rightarrow \infty$

## Simple 1-D example

Gaussian random walk on  $X$ -axis, s.d.  $\sigma_x$ , sensor s.d.  $\sigma_z$

$$\mu_{t+1} = \frac{(\sigma_t^2 + \sigma_x^2)z_{t+1} + \sigma_z^2\mu_t}{\sigma_t^2 + \sigma_x^2 + \sigma_z^2}$$

$$\sigma_{t+1}^2 = \frac{(\sigma_t^2 + \sigma_x^2)\sigma_z^2}{\sigma_t^2 + \sigma_x^2 + \sigma_z^2}$$



# General Kalman update

Transition and sensor models:

$$\begin{aligned}P(\mathbf{x}_{t+1}|\mathbf{x}_t) &= N(\mathbf{F}\mathbf{x}_t, \mathbf{\Sigma}_x)(\mathbf{x}_{t+1}) \\P(\mathbf{z}_t|\mathbf{x}_t) &= N(\mathbf{H}\mathbf{x}_t, \mathbf{\Sigma}_z)(\mathbf{z}_t)\end{aligned}$$

$\mathbf{F}$  is the matrix for the transition;  $\mathbf{\Sigma}_x$  the transition noise covariance

$\mathbf{H}$  is the matrix for the sensors;  $\mathbf{\Sigma}_z$  the sensor noise covariance

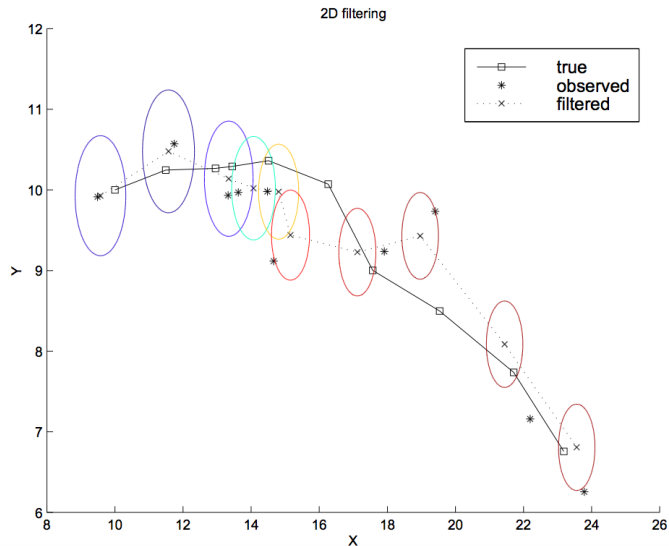
Filter computes the following update:

$$\begin{aligned}\boldsymbol{\mu}_{t+1} &= \mathbf{F}\boldsymbol{\mu}_t + \mathbf{K}_{t+1}(\mathbf{z}_{t+1} - \mathbf{H}\mathbf{F}\boldsymbol{\mu}_t) \\ \mathbf{\Sigma}_{t+1} &= (\mathbf{I} - \mathbf{K}_{t+1})(\mathbf{F}\mathbf{\Sigma}_t\mathbf{F}^\top + \mathbf{\Sigma}_x)\end{aligned}$$

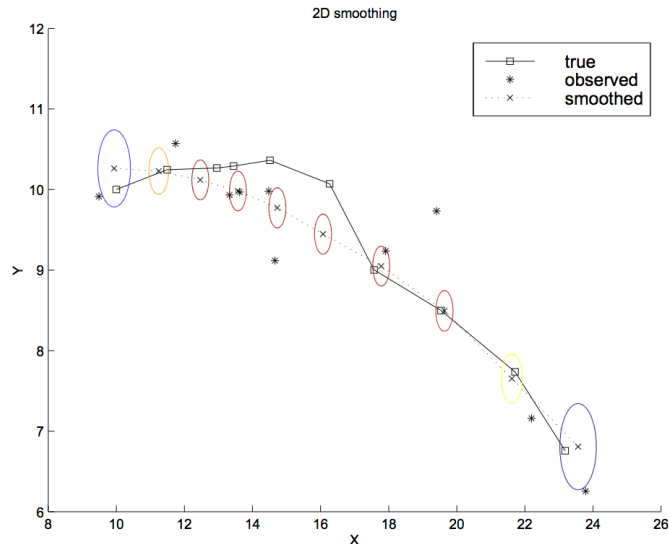
where  $\mathbf{K}_{t+1} = (\mathbf{F}\mathbf{\Sigma}_t\mathbf{F}^\top + \mathbf{\Sigma}_x)\mathbf{H}^\top (\mathbf{H}(\mathbf{F}\mathbf{\Sigma}_t\mathbf{F}^\top + \mathbf{\Sigma}_x)\mathbf{H}^\top + \mathbf{\Sigma}_z)^{-1}$   
is the **Kalman gain matrix**

$\mathbf{\Sigma}_t$  and  $\mathbf{K}_t$  are independent of observation sequence, so compute offline

## 2-D tracking example: filtering



## 2-D tracking example: smoothing



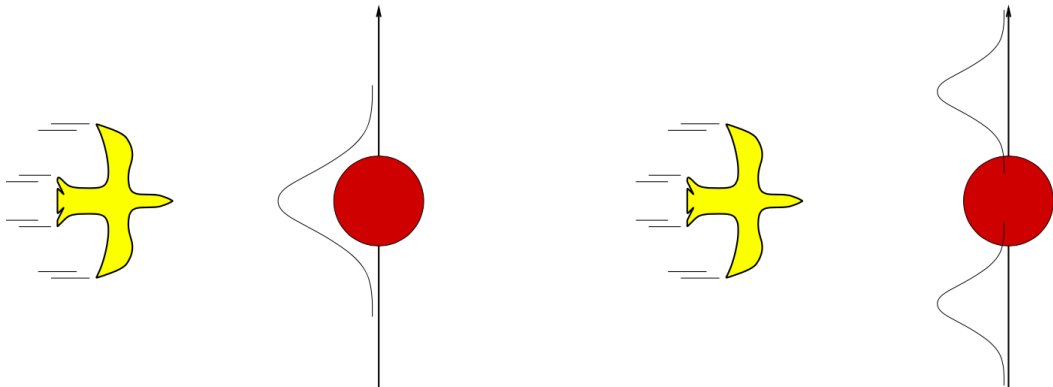


## Where it breaks

Cannot be applied if the transition model is nonlinear

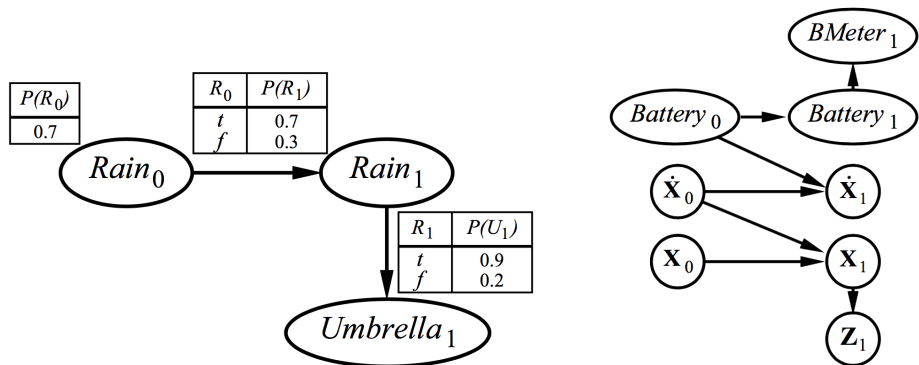
Extended Kalman Filter models transition as **locally linear** around  $\mathbf{x}_t = \boldsymbol{\mu}_t$

Fails if system is locally unsmooth



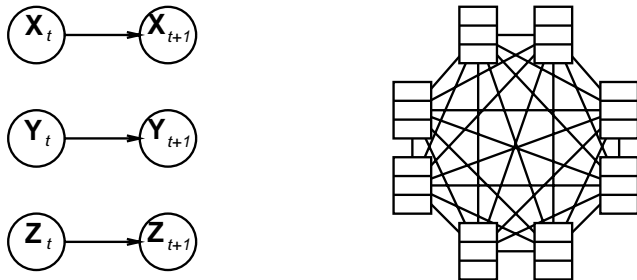
# Dynamic Bayesian networks

$\mathbf{X}_t, \mathbf{E}_t$  contain arbitrarily many variables in a replicated Bayes net



## DBNs vs. HMMs

$\mathbf{X}_t, \mathbf{E}_t$  contain arbitrarily many variables in a replicated Bayes net  
Every HMM is a single-variable DBN; every discrete DBN is an HMM



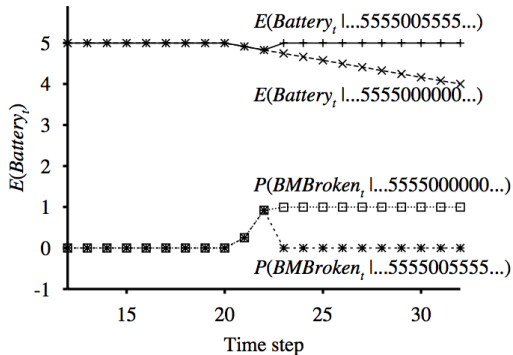
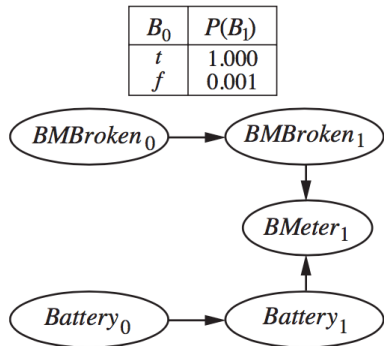
Sparse dependencies  $\Rightarrow$  exponentially fewer parameters;

e.g., 20 state variables, three parents each

DBN has  $20 \times 2^3 = 160$  parameters, HMM has  $2^{20} \times 2^{20} \approx 10^{12}$

# DBNs vs. Kalman filters

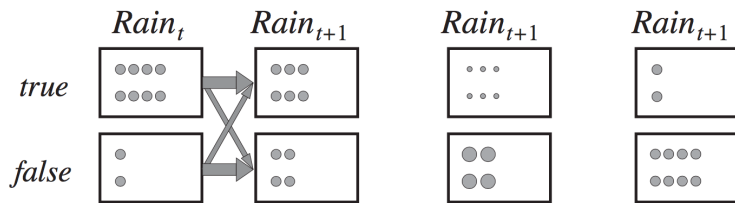
Every Kalman filter model is a DBN, but few DBNs are KFs;  
real world requires non-Gaussian posteriors



# Particle filtering

Basic idea: ensure that the population of samples (“particles”) tracks the high-likelihood regions of the state-space

Replicate particles proportional to likelihood for  $\mathbf{e}_t$



Widely used for tracking nonlinear systems, esp. in vision

Also used for simultaneous localization and mapping in mobile robots  
 $10^5$ -dimensional state space

# Summary

Temporal models use state and sensor variables replicated over time

Markov assumptions and stationarity assumption, so we need

- transition model  $\mathbf{P}(\mathbf{X}_t | \mathbf{X}_{t-1})$
- sensor model  $\mathbf{P}(\mathbf{E}_t | \mathbf{X}_t)$

Tasks are filtering, prediction, smoothing, most likely sequence;

**all done recursively with constant cost per time step**

Hidden Markov models have a single discrete state variable; used for speech recognition

Kalman filters allow  $n$  state variables, linear Gaussian,  $O(n^3)$  update

Dynamic Bayes nets subsume HMMs, Kalman filters; exact update intractable

Particle filtering is a good approximate filtering algorithm for DBNs